# A Virtual Closed Loop Method for Closed Loop Identification $\star$

Juan C. Agüero<sup>a</sup>, Graham C. Goodwin<sup>a</sup>, Paul M. J. Van den Hof<sup>b</sup>

<sup>a</sup>Centre for Complex Dynamic Systems and Control (CDSC) School of Electrical Engineering and Computer Science University of Newcastle, Australia

> <sup>b</sup>Delft Center for Systems and Control Delft University of Technology The Netherlands

#### Abstract

Indirect methods for identification of linear plant models on the basis of closed-loop data are based on the use of (reconstructed) input signals that are uncorrelated with the noise. This generally requires exact (linear) controller knowledge. On the other hand, direct identification requires exact plant and noise modelling (system in the model set) in order to achieve accurate results, although the controller can be nonlinear. In this paper, a generalized approach to closed-loop identification is presented that includes both methods as special cases and which allows novel combined methods to be generated. Besides providing robustness with respect to inexact controller knowledge, the method does not rely on linearity of the controller nor on exact noise modeling. The generalization is obtained by balancing input-noise decorrelation against noise whitening in a user-chosen flexible fashion. To this end, a user-chosen virtual controller is used to parametrize the plant model, thereby generalizing the dual-Youla method to cases where knowledge of the controller is inexact. Asymptotic bias and variance results are presented for the method. Also, the benefits of the approach are demonstrated via simulation studies.

Key words: System identification, Closed Loop Identification.

# 1 Introduction

Identification of dynamic systems operating in the presence of feedback has received considerable attention in the system identification literature (see e.g. [16,30,33,8,23]). In many situations, there exist strong economic and/or safety reasons for requiring process data to be collected under closed loop conditions. Even open-loop stable plants are often subject to non-stationary disturbances and long term drift that favour a closed loop experiment for data collection. Additionally, it has been found that, for several model applications (as e.g. model-based control design) and experimental constraints (e.g. output power constraints), a closed-loop identification experiment is often the optimal experimental setup, see e.g. [34,21,10,20,4].

Unfortunately, handling closed-loop data in identification

leads to additional difficulties, see e.g. [33,8].

In the prediction error framework [23], the two principle methods for closed loop identification can be characterized as follows:

- Direct identification: Here the plant is identified directly on the basis of plant input and output data taken from within the closed-loop. The presence of feedback is ignored. Consistent model estimates can be identified under the condition that the noise dynamics are modelled exactly. Exact knowledge of the controller is not necessary and the controller may be nonlinear.
- Indirect identification: Here a plant object<sup>1</sup> is identified (usually the complementary sensitivity) between the reference input and plant output signals. Subsequently an equivalent plant model is retrieved from the identified object. Consistent plant models can be identified under the condition that the controller is linear and exactly known. This holds for several variants of the indirect method, including the dual-Youla approach [17,28,33],

<sup>\*</sup> This paper was not presented at any IFAC meetings. Corresponding author: Juan C. Agüero. Tel. +61-2-49216351. Fax +61-2-49601712.

Email addresses: juan.aguero@newcastle.edu.au (Juan C. Agüero), graham.goodwin@newcastle.edu.au (Graham C. Goodwin), p.m.j.vandenhof@tudelft.nl (Paul M. J. Van den Hof).

<sup>&</sup>lt;sup>1</sup> Here and in the sequel, we use the term "*plant object*" to describe a transfer function that depends on the (open-loop) plant.

the method based on tailor-made parametrization [35], and bias-elimination least-squares methods (BELS) [39].

A third category, called joint input-output methods (see e.g. [30]), can be considered as an indirect method in the context of the current paper.

In practice, one would ideally like to have identification tools that combine the advantageous properties of both direct and indirect methods. In particular, one would like to be able to handle the following situations:

- When a (slightly) nonlinear controller is present in the loop, e.g. a linear controller that saturates regularly, or that is not exactly known;
- When the noise disturbances are non-stationary or cannot be modelled exactly by stationary white noise filtered through a linear time-invariant system.

In the above situations, neither direct nor indirect methods provide consistent model estimates and the number of alternatives is very limited. The Projection Method [9] was proposed to deal with non-linear controllers, by identifying non-causal FIR models to approximate non-linear sensitivity functions. Alternatively, Instrumental Variable methods [31,13] can handle non-linear controllers and yield consistent plant models irrespective of noise under-modeling.

In this paper we develop a novel approach that takes a more generalized perspective. Realizing that all indirect methods use controller knowledge to exactly decorrelate the identification input signal from the noise, we will focus on this decorrelation and develop a generalized approach that is robust against controller nonlinearity and inexact controller knowledge.

A central issue in our development is the choice of a virtual controller [2,1,14] which approximates the real (possibly nonlinear) controller. This virtual controller will be deployed only for input signal construction to be used in identification. The more accurate the virtual controller, the less noise correlation will be present in the identification input. As a result, the bias due to having an inexact noise model will be reduced. In this way our method generalizes both the direct and indirect method of closed-loop identification. Our approach leads to a sliding mechanism between these two extremes. The user can make an appropriate choice depending on his/her faith in either the quality of the noise model, or the available knowledge and linearity of the controller.

The current paper completes and generalizes the analysis originally presented in [5], see also [2] and [14].

The remainder of the paper is organized as follows: In Section 2 we introduce closed loop identification in a general non-linear setting. In Section 3 we describe the new approach for identification of closed loop systems. In Section 4 we show how the virtual closed loop (VCL) method generalizes known schemes for closed loop identification. In Section 5 we analyze the spectra of signals appearing in the VCL method. In Section 6 we show how the choice of parameters in the virtual closed loop method affects the asymptotic bias and the estimation accuracy of the identification for systems operating in closed loop. We also analyze the accuracy of the estimates provided by VCL. In Section 7 we provide general guidelines to design filters necessary to implement the identification procedure using VCL. In Section 8 we illustrate how to use VCL to identify a simple system. Finally in section 9 we draw conclusions.

## 2 Closed-loop identification setup

We consider a data generating system S:

$$y_t = G_o(q)u_t + v_t \tag{1}$$
$$v_t = H_o(q)w_t$$

where q is the forward-shift operator,  $y_t$ ,  $u_t$ , and  $w_t$  are the output, input and noise respectively,  $G_o(q)$ , and  $H_o(q)$  are linear transfer functions, with  $H_o$  stable, stably invertible and monic (i.e.  $\lim_{|z|\to\infty} H_o(z) = 1$ ). The noise  $w_t$  is assumed to be zero mean Gaussian white noise with variance  $\sigma_w^2$ . The system is assumed to operate in a stabilized closed loop (see Figure 1). In the case that the controller is linear the input signal satisfies:

$$u_t = C(r_t - y_t),\tag{2}$$

for non-linear controller the relationship in 2 has to be understood as a nonlinear dynamic map between the tracking error  $r_t - y_t$  and the input signal  $u_t$ .

Here,  $r_t$  is an external reference signal. Throughout the paper, we will not restrict the controller, C, to be linear . However, at times, it will be convenient to consider the linear case so that we can relate our work to earlier literature. In these cases, we will use the notation  $C_l$  to denote C. In addition, we will, at other times, wish to consider a linear controller which is "close" (in some sense) to a non-linear controller. In this case, we will use the notation  $C_l^a$ .



Fig. 1. Closed-loop system configuration

In order to have a well-defined closed-loop, it is further assumed that either C or  $G_o$  contains, at least, a one step

time delay.

Spectral densities of signals are denoted by

$$\Phi_{uw}(\omega) = \sum_{\tau = -\infty}^{\infty} R_{uw}(\tau) e^{-j\omega\tau}$$

with <sup>2</sup>  $R_{uw}(\tau) := \overline{\mathbf{E}} \{ u_t w_{t-\tau} \}$  and  $\Phi_u = \Phi_{uu}$ .

In order to be able to define signal spectra and cross-spectra we assume that the following assumption holds [23,8]:

**Assumption 1** The signals  $w_t$ ,  $r_t$ ,  $u_t$ ,  $y_t$  in the closed loop system defined by (1)-(2) are jointly quasi-stationary.  $\nabla \nabla \nabla$ 

Our goal is the identification of a (consistent) plant model for  $G_o$  on the basis of closed-loop data. There are key differences between direct and indirect methods of identification. The different conditions for arriving at consistent estimates of  $G_o$  are listed in Table 1.

	direct	indirect
Exact LTI noise model	yes <sup>3</sup>	no
C linear	no	yes
C exactly known	no	yes

Table 1

Required conditions for direct and indirect identification to provide consistent plant model estimates.

The clear distinction between the two situations raises the question as to whether one can combine the two approaches in a generalized method that is robust with respect to all three conditions; i.e. a method that is robust with respect to slight deviations from the assumptions of having exact controller knowledge, controller linearity, or exact LTI noise models. Such a method is developed in the next section.

#### **3** A generalized approach to closed-loop identification

#### 3.1 Introduction

In our generalized approach we consider the setup of Figure 2, where linear, causal and stable filters  $F_1 \cdots F_4$  are introduced to generate signals

$$x_t = F_1(q)u_t + F_2(q)y_t$$
(3)

$$z_t = F_3(q)u_t + F_4(q)y_t.$$
 (4)

<sup>2</sup> Here and in the sequel, we use the operator  $\bar{E}\{(\cdot)\} = \lim_{N \to \infty} \frac{1}{N} \sum_{t=1}^{N} E\{(\cdot)\}$  where  $E\{(\cdot)\}$  is the expected value operator

value operator.

<sup>3</sup> Notice that for estimation techniques such as instrumental variables, it is possible to identify the system  $G_o$  using direct identification without modelling the noise transfer function.

The signals  $x_t$  and  $z_t$  will be used as input and output signals in a generalized identification scheme.

Notice that Assumption 1 and the relationships (3)-(4) imply that not only are all the signals in the closed loop (1)-(2) quasi-stationary but also the signals  $x_t$  and  $z_t$ .



Fig. 2. Generalized scheme for closed-loop identification from input  $x_t$  to output  $z_t$ .

The proposed identification approach amounts to identifying a plant-related object (parameterized by a vector  $\theta$  that defines the model for the system) through a linear transfer function between the signals  $x_t$  and  $z_t$ , by applying a model structure:

$$\varepsilon_t(\theta) = K(q,\theta)^{-1}[z(t) - R(q,\theta)x(t)]$$
(5)

leading to estimates  $\hat{R} = R(q, \hat{\theta}_N)$  and  $\hat{K} = K(q, \hat{\theta}_N)$ , and subsequently to derive an equivalent plant model  $\hat{G}$  from  $\hat{R}$  by applying the principle of tailor-made parametrization. (More details will be provided in sub-section 3.4)

It is apparent that the identification results will depend on the choice of the filters  $F_1 \cdots F_4$ . For example, by appropriately choosing  $F_1$  and  $F_2$  one can tune the presence of noise  $w_t$  in the generalized input signal  $x_t$ , and thereby one can influence the resulting bias.

#### 3.2 The virtual closed-loop

Under, the additional condition that  $F_1$  is stably invertible (Note that this is not required in the method), the closed-loop diagram of Figure (2) can be redrawn as shown in Figure (3). This alternative view shows how the filters  $F_1$  and  $F_2$ are used to construct the generalized reference input  $x_t$  via  $x_t = F_1u_t + F_2y_t$  (where  $F_2/F_1$  acts as a virtual controller that compensates for the original controller in the construction of  $x_t$ ). Additionally, the resulting transfer function between  $x_t$  and  $z_t$  contains a (virtual) closed-loop plant object, where again the same controller  $F_2/F_1$  is involved. Since the (linear) filter  $F_2/F_1$  can be chosen freely by the user, it



Fig. 3. Virtual Closed Loop.

does not have any direct relation to the implemented controller C. Therefore the scheme is referred to by the name "Virtual Closed Loop". We also define the virtual controller as  $\overline{C}(q) = F_2(q)/F_1(q)$ . Note also that stability is not an issue since we already know that all signals are bounded (in a suitable statistical sense). Moreover, in the case that the "true" controller is linear and equal to the virtual controller  $(C_l = \overline{C} = F_2/F_1)$  then  $x_t = F_2r_t$  is uncorrelated with the noise  $w_t$ . In the remainder of the paper, we will explore the implications of the configuration of Figure 2 in system identification.

#### 3.3 System equations and filter conditions

In order to analyse the setup, we use equations (1), (3) and (4) to obtain the following set of equations describing the virtual closed loop system:

$$\begin{bmatrix} 1 & -F_3 & -F_4 \\ 0 & F_1 & F_2 \\ 0 & -G_o & 1 \end{bmatrix} \begin{bmatrix} z_t \\ u_t \\ y_t \end{bmatrix} = \begin{bmatrix} 0 \\ x_t \\ v_t \end{bmatrix}$$
(6)

Solving for z, u and y we have the following:

$$z_t = \frac{F_3 + F_4 G_o}{F_1 + F_2 G_o} x_t + \frac{F_1 F_4 - F_2 F_3}{F_1 + F_2 G_o} H_o w_t$$
(7)

$$u_t = \frac{1}{F_1 + F_2 G_o} x_t - \frac{F_2}{F_1 + F_2 G_o} H_o w_t$$
(8)

$$y_t = \frac{G_o}{F_1 + F_2 G_o} x_t + \frac{F_1}{F_1 + F_2 G_o} H_o w_t$$
(9)

We then write

$$z_t = R_o x_t + K_o w_t \tag{10}$$

with

$$R_o = \frac{F_3 + F_4 G_o}{F_1 + F_2 G_o}, \quad K_o = \frac{F_1 F_4 - F_2 F_3}{F_1 + F_2 G_o} H_o$$
(11)

In order for  $R_o$  and  $K_o$  to satisfy the usual conditions for applying prediction error identification methods the filters  $F_1 \cdots F_4$  have to satisfy certain regularity conditions.

**Assumption 2** We assume that the filters  $\{F_i\}_{i=1...4}$  satisfy the following:

- (1) F<sub>1</sub> is biproper;
  (2) F<sub>2</sub>G<sub>o</sub> is strictly proper;
- (3)  $G_o$  is stabilized by the controller  $F_2/F_1$ ;
- (4)  $M := F_1F_4 F_2F_3$  is stably invertible.

 $\nabla \nabla \nabla$ 

**Proposition 3** The transfer functions  $R_o$  and  $K_o$  are causal and stable under the conditions in Assumption 2.

# Proof

Conditions (1) and (2) together with the causality of all filters  $F_1, \dots F_4$  guarantee that  $R_o$  and  $K_o$  are causal (i.e.  $\lim_{z\to\infty} R_o(z)$ , and  $\lim_{z\to\infty} K_o(z)$  are constants).

In order to show stability of  $R_o$  and  $K_o$ , we express the filters  $\{F_i\}_{i=1..4}$  and the plant  $G_o$  as a coprime polynomial factorizations:  $F_i = N_i D_i^{-1}$ ,  $i = 1, \dots 4$ , and  $G_o = B_o/A_o$ . Then

$$R_o = \frac{D_1 D_2}{D_3 D_4} \frac{(N_3 D_4 A_o + N_4 D_3 B_o)}{(N_1 D_2 A_o + N_2 D_1 B_o)}$$
(12)

Stability of  $R_o$  and  $K_o$  follows if  $N_1D_2A_o + N_2D_1B_o$  has all zeros within the unit circle (see e.g. [15, page 127]). This is guaranteed when  $G_o$  is stabilized by the controller  $N_2D_1/N_1D_2 = F_2/F_1$ , see condition (3).

**Remark 4** Notice that since the noise  $w_t$  is Gaussian distributed, then without loss of generality we assume that  $K_o$  is minimum-phase. In addition, in the estimation procedure we assume that K is also minimum-phase.  $\nabla \nabla \nabla$ 

#### 3.4 Identification setup

The virtual closed-loop system will be identified by applying a direct identification method to the virtual-closed-loop system (10). On the basis of the generalized input signal  $x_t$ , and the generalized output signal  $z_t$ , we identify a model in a Box-Jenkins type model structure, with a prediction error

$$\varepsilon_t = K(q, \eta)^{-1} \left[ z_t - R(q, \rho) x_t \right]$$
(13)

where

$$R(q,\rho) = \frac{F_3 + F_4 G(q,\rho)}{F_1 + F_2 G(q,\rho)}.$$
(14)

represents the family of models used to estimate  $R_o$  and  $K(q,\eta)$  represents the family of models used to estimate  $K_o$ .

Preprint sub<sup>4</sup>itted to Automatica Received February 15, 2011 15:47:37 PST The parametrization of  $R(q, \rho)$  is a tailor-made parametrization in which the parameters of the plant model G are used to parametrize the virtual closed-loop plant  $R(q, \rho)$  [8,35]. In the parametrization (13), the plant and noise models are parametrized independently. The parameters  $\rho$ , and  $\eta$  are estimated by minimizing a quadratic criterion:

$$V_N = \frac{1}{N} \sum_{t=1}^N \varepsilon_t^2.$$
 (15)

The following condition on the filters  $\{F_i\}_{i=1..4}$ , the true system  $G_o$ , and the class of models G is assumed to hold:

$$\lim_{|z| \to \infty} \frac{(F_1 F_4 - F_2 F_3)}{(F_1 + F_2 G_o)} = \lim_{|z| \to \infty} \frac{(F_1 F_4 - F_2 F_3)}{(F_1 + F_2 G)} = 1$$

This condition implies that  $K_o$  and K are monic.

Notice that once estimates  $\hat{G}$ ,  $\hat{K}$  for  $G_o$  and  $K_o$  have been obtained, one can define an estimate for  $H_o$  as follows:

$$\hat{H} := \frac{F_1 + F_2 \hat{G}}{F_1 F_4 - F_2 F_3} \hat{K}$$
(16)

We make the following assumption:

**Assumption 5** The vector of parameters  $\theta = [\rho^T \eta^T]^T$ , the input  $(x_t)$ , noise  $(w_t)$  and reference  $(r_t)$  satisfy regularity conditions such that the solution,  $\hat{\theta}_N$  of the optimization problem in (15) converges (a.s.) to  $\theta_*$ .  $\nabla \nabla \nabla$ 

Assumption 5 is necessary for asymptotic statistical analysis to hold (see [22] and [37] for details). Sufficient conditions on the true system, signals and the parametrized family of models such that Assumption 5 holds have to be obtained for every particular case. Moreover, the asymptotic statistical analysis presented in [22] also holds when the noise  $w_t$  is a i.i.d. sequence not necessarily Gaussian distributed.

In the subsequent analysis we will analyze the impact of the following two issues:

- (1)  $x_t$  is not, in general, an exogenous signal but is potentially correlated with the noise  $w_t$ .
- (2) The class of models used for K(q, η) may not include the true noise model K<sub>o</sub> e.g. we might decide to use a fixed noise model K ≠ K<sub>o</sub>.

# 4 Particular cases and general properties

We will first establish that the Virtual Closed Loop method generalizes known methods for closed loop identification.

• Direct identification (see e.g. [23]) is obtained by the choice  $F_1 = F_4 = 1$ , and  $F_2 = F_3 = 0$ . This results in

$$x = u, \ z = y, \ R_o = G_o, \ K_o = H_o$$

• Traditional indirect identification ([30]) is obtained when the controller is linear and the choice  ${}^4$   $F_1 = C_l^{-1}$ ,  $F_2 = F_4 = 1$ ,  $F_3 = 0$  is made where  $C_l$  is the (assumed known and linear) true controller. This results in

$$x = r, \ z = y, \ R_o = \frac{G_o C_l}{1 + G_o C_l}, \ K_o = \frac{1}{1 + G_o C_l} H_o.$$

Notice that it is necessary to incorporate an extra signal in the traditional indirect identification (the reference x = r). However, if the "true" controller is linear, then it is possible to reconstruct the reference from the input/output signals.

• In the Dual Youla method ([17,28,34]) the plant model parametrization is based on an auxiliary model  $G_x$  of  $G_o$  with rational coprime factorization  $N_x/D_x$  that is stabilized by the present (assumed known and linear) controller  $C_l$  with rational coprime factorization  $N_c/D_c$ . This method is obtained by choosing

$$F_1 = D_c/M; \ F_2 = N_c/M; \ F_3 = -N_x/M; \ F_4 = D_x/M$$

with  $M = N_c N_x + D_c D_x$ .

• The "whitening procedure" (see e.g. [7,23]) is obtained by the choice  $F_1 = F_4 = F$  and  $F_2 = F_3 = 0$ . In this case we have

$$x = Fu, \ z = Fy, \ R_o = G_o, \ K_0 = FH_o.$$

Note that if  $F \approx H_o^{-1}$ , then we might consider using a fixed filter K = 1 in the estimates.

The different methods and the corresponding choices for the filters  $F_1 \cdots F_4$  are listed in Table 2. The identification objects and input and output signals are collected in Table 3. Note that the indirect and Dual-Youla methods require the controller C to be linear.

	$F_1$	$F_2$	$F_3$	$F_4$
Direct	1	0	0	1
Indirect	$C_l^{-1}$	1	0	1
Dual-Youla (DY)	$D_c/M$	$N_c/M$	$-N_x/M$	$D_x/M$

Table 2

Particular choice for the filters  $F_1, \dots, F_4$  leads to specific closedloop identification methods.  $M = N_c N_x + D_c D_x$ .

<sup>&</sup>lt;sup>4</sup> If  $C_l^{-1}$  is non-causal and the reference signal is available, then one can directly define  $x_t = r_t$ .

	input x	output z	Ro	Ko
Direct	u	y	$G_o$	$H_o$
Indirect	$C_l^{-1}u + y$	y	$\frac{C_l G_o}{1 + C_l G_o}$	$\frac{H_o}{1+C_lG_o}$
Dual-Youla (DY)	$\frac{D_x^{-1}}{1+C_l G_x}(u+Cy)$	$\frac{D_c^{-1}}{1+C_l G_x}(y-G_x u)$	$\frac{(G_o - G_x)D_x}{D_c(1 + C_lG_o)}$	$\frac{D_c^{-1}H_o}{1+C_lG_o}$
VCL	$F_1u + F_2y$	$F_3u + F_4y$	$\frac{F_3 + F_4 G_o}{F_1 + F_2 G_o}$	$\left  \frac{F_1 F_4 - F_2 F_3}{F_1 + F_2 G_o} H_o \right $

Table 3

Overview of input / output signals and of objects of identification for closed-loop identification methods.

The following observations reflect some of the main properties associated with selecting the filters  $F_1 \cdots F_4$  in our method:

- If the model sets are flexible enough to capture the real plant and noise dynamics of  $R_o$  and  $K_o$  respectively, then all methods provide consistent estimates of  $G_o$  and  $H_o$ .
- If the model sets for R<sub>o</sub> are chosen flexible enough to represent the real plant dynamics, (and no statement is made with respect to model sets for K<sub>o</sub>), then the plant estimates Ĝ will contain an asymptotic bias that is determined by Φ<sub>xw</sub>. This bias is zero when Φ<sub>xw</sub> = 0.
- The input signal  $x_t$  for identification is uncorrelated with the noise  $w_t$ , i.e.  $\Phi_{xw} = 0$ , if the controller is linear and the virtual controller is chosen as  $\overline{C} := F_2/F_1 = C_l$ .
- If the auxiliary model  $F_3/F_4$  is stabilized by the virtual controller  $\overline{C}$ , then any identified model  $\hat{R}$  of  $R_o$  that is stable will, through (14), correspond to an equivalent plant model  $\hat{G}$  that is stabilized by the virtual controller  $\overline{C}$ .
- The virtual closed-loop method incorporates a generalized Dual-Youla method, where the controller that is used for the plant parametrization is not necessarily chosen equal to the present (possibly nonlinear) controller C, but is a user-chosen linear approximation thereof in the form of the virtual controller  $\overline{C}$ .

In [8, lemma 3], it is established that traditional indirect identification can be thought as a direct identification method where the noise model is parametrized in terms of the open loop process G and the true controller, C. The analysis presented in [8, lemma 3] assumes that the true controller is linear and exactly known. We next, generalize this result for the case of the VCL method.

**Lemma 6** *VCL identification is equivalent to direct identification with the following prediction model:* 

$$y_t = G(q,\rho)u_t + \bar{H}(q,\rho,\eta)\epsilon_t \tag{17}$$

where the noise model is given by:

$$\bar{H}(q,\rho,\eta) = K(q,\eta) \frac{F_1 + F_2 G(q,\rho)}{F_1 F_4 - F_2 F_3}$$
(18)

**Proof:** In the VCL method the prediction error is given by:

$$\epsilon_t = \frac{1}{K} \left[ z_t - \frac{F_3 + F_4 G}{F_1 + F_2 G} x_t \right] \tag{19}$$

Using the input-output relationship and re-arranging terms we obtain:

$$\epsilon_t = \frac{1}{K} \frac{M}{F_1 + F_2 G} [y_t - Gu_t] \tag{20}$$

$$M = F_1 F_4 - F_2 F_3 \tag{21}$$

*The result follows since the same prediction error is obtained from direct identification with noise model*  $\overline{H}(q, \rho, \eta)$ .  $\Box$ 

The previous lemma shows that VCL is equivalent to shaping the noise model for the system to be identified. This lemma also shows that most indirect identification methods can be considered as a modified version of direct identification.

Notice that the set of poles of the extended noise model  $\overline{H}$  contains the poles of G. This property also appears in ARMAX and ARX models and is actually the key enabling tool that allows one to identify unstable processes  $G_o$ .

Even though, VCL can be understood as a special case of direct identification, the key point is that a systematic procedure exists to modify the effective noise model in order to reduce the bias due to under-modeling and due to signal correlation arising from the closed loop nature of the data. Note that, in the usual direct identification for Box-Jenkins (BJ) models, the noise model and the plant are independently parametrized. This means, that it is not possible, in general, to identify unstable systems using a Box-Jenkins parametrization. By way of contrast, in VCL, even though we are using BJ models, the identification procedure when viewed in the direct identification setting is not a Box-Jenkins model, but has a very particular structure.

### 5 Signal spectra analysis for the VCL

#### 5.1 Preliminary definitions and assumptions

**Definition 7** [29, page 197] If the transfer function X(z) is given by:

$$X(z) = \dots + x_{-1}z^{1} + x_{0} + x_{1}z^{-1} + x_{2}z^{-2} + \dots$$
 (22)

where  $z \in A \subset \mathbb{C}$ , and A includes the unit circle, then the causal part is given by:

$$[X(z)]_{+} := x_0 + x_1 z^{-1} + x_2 z^{-2} + \dots$$
 (23)

and the anti-causal part by:

$$[X(z)]_{-} := X(z) - [X(z)]_{+} = x_{-1}z^{1} + x_{-2}z^{2} + \dots$$
(24)
$$\nabla \nabla \nabla$$

We assume some extra conditions in order to have a well defined closed loop system.

Assumption 8 One of the following two conditions holds:

- the true plant,  $G_o$ , and its model, G, are strictly causal,
- the true controller, C, and the virtual controller, C are strictly causal. ∇∇∇

Note that Assumption 8 imposes a constraint in the class of models G and K.

### 5.2 Cross-spectrum between $x_t$ and $w_t$

A common problem in the identification of closed loop systems is that the input to the system to be identified is correlated with the noise [23]. We next analyze the impact that the choice of the different filters  $\{F_i\}_{i=1..4}$  has on the cross-spectrum  $\Phi_{xw}$  given by:

$$\Phi_{xw} = (F_1 + F_2 G_o) \Phi_{uw} + F_2 H_o \sigma_w^2$$
(25)

The following result provides a condition such that the cross-spectrum  $\Phi_{xw}$  is identically zero.

**Lemma 9** The cross-spectrum between x and w vanishes if

$$F_1(1+G_o\overline{C})\Phi_{uw} + F_2H_o\sigma_w^2 = 0$$
(26)

Moreover (26) holds if the virtual controller is given by:

$$\overline{C} = -\frac{\Phi_{uw}}{G_o \Phi_{uw} + H_o \sigma_w^2} \tag{27}$$

**Proof:** Immediate from equations (25) and  $\overline{C} = \frac{F_2}{F_1}$ ,  $\overline{S}_o = \frac{1}{1+G_o\overline{C}}$ .

An implication of Lemma 9 is that it is possible to reduce the correlation between  $x_t$  and  $w_t$  by adjusting  $\overline{C}$  irrespective of the linearity of the true controller.

We next specialize to the case when the true controller has a linear approximation which we denote  $C_l^a$ .

**Corollary 10** If the input of the real system is given by the following relationship:

$$u_t = C_l^a(q)(r_t - y_t) + \xi_t$$
(28)

where  $C_l^a(q)$  is a linear controller, and  $\xi_t$  is a quasistationary signal that might depend on  $r_t$ ,  $y_t$  and their past values, then condition (27) can be re-written as:

$$\overline{C} = \overline{\beta}C_l^a + (1 - \overline{\beta})(-G_o^{-1}) \tag{29}$$

where

$$\bar{\beta} = \frac{1}{1 + \frac{G_o}{H_o} \frac{\Phi_{\xi w}}{\sigma_w^2}} \tag{30}$$

**Proof:** *Re-writing equation* (27), *and using the relationship between input and output signals we observe that the input signal is given by:* 

$$u_t = C_l^a S_o^a r_t + S_o^a \xi_t - C_l H_o S_o^a w_t$$
(31)

where  $S_o^a$  is the sensitivity function calculated using  $C_l^a$ . Finally, considering that  $\Phi_{xw} = \frac{F_1}{S_o} \Phi_{uw} + F_2 H_o \sigma_w^2$ , and re-arranging terms we obtain the result.  $\Box$ 

Notice that, if the true controller is non-linear, then  $C_l^a(q)$  in (28) represents a linear approximation of the true controller, and  $\xi_t$  captures the remaining terms due to non-linearities.

From the previous corollary we see that, if the controller is slightly non-linear ( $\Phi_{\xi w}$  small), then  $\bar{\beta} \approx 1$ , and thus one can reduce  $\Phi_{xw}$  by choosing the virtual controller as a linear approximation of the true non-linear controller ( $\bar{C} \approx C_l^a$ ).

#### 6 Bias and accuracy analysis for the VCL method

Bias and variance are fundamental concepts to asses the performance of estimators [32]. We next, analyze the bias and variance of the estimators in the VCL setup.

It is well known that the cost function in direct identification using PEM is asymptotically (in the number of data points) given by  $^5$  [23]:

$$V_{N} \rightarrow \frac{1}{2\pi} \int_{\pi}^{\pi} \Phi_{\epsilon}$$
  
= $\sigma_{w}^{2} + \frac{1}{2\pi} \int_{\pi}^{\pi} \frac{\Phi_{u}}{|H|^{2}} \left| (G_{o} - G) + (H_{o} - H) \frac{\Phi_{wu}}{\Phi_{u}} \right|^{2}$   
+ $\frac{1}{2\pi} \int_{\pi}^{\pi} \frac{|H_{o} - H|^{2}}{|H|^{2}} \left( \sigma_{w}^{2} - \frac{|\Phi_{wu}|^{2}}{\Phi_{u}} \right)$  (32)

The expression in (32) is valid for different system parametrizations and holds irrespective of the conditions under which the real system is operating (i.e. open or closed loop). If the system is operating in open loop and one uses a Box-Jenkins models structure (i.e.  $G(q, \theta) = G(q, \rho)$  and  $H(q, \theta) = H(q, \eta)$ , and  $\theta = [\rho^T \eta^T]^T$ ), then we have that the corresponding optimization problem to determine  $\rho$  is (asymptotically in the number of data points) given by:

$$\hat{\rho} = \arg\min_{\rho} \frac{1}{2\pi} \int_{\pi}^{\pi} \frac{\Phi_u}{|H_*|^2} |G_o - G|^2$$
(33)

where  $H_*(q) = H(q, \eta_*)$  and  $\eta_*$  is the value for  $\eta$  that optimizes the asymptotic cost function (32). Thus, in this particular case, we have that, if the model class for G contains  $G_o$ , then one obtains a consistent estimate for  $G_o$  irrespective of possible under-modelling in the noise transfer function  $H_o$  (i.e.  $H_*(q) \neq H_o(q)$ ). This analysis illustrates the benefits of using Box-Jenkins models for systems operating in open loop.

For systems operating in closed loop, the term  $B_G = \frac{\Phi_{wu}}{\Phi_u}(H_o - H_*)$  is usually called the bias-pull for the estimates of  $G_o$  [8]. We next present a definition of this term that will be useful for the analysis in the sequel.

**Definition 11** Consider a given true system  $G_o(q)$ , and a given cost function  $V_N(\theta)$  that converges to  $V_{\infty}(\theta)$  as  $N \to \infty$ . The bias-pull is the bias that would occur in an estimated G-model if it were parametrized non-parametrically (i.e. allowing unlimited orders and non-causal dynamics) and independent of the noise model.  $\nabla \nabla \nabla$ 

The importance of the bias-pull concept is that, for BJ models, it defines the limit point where an estimate of  $G_o(q)$ obtained by solving an optimization problem converges, but limited to the freedom available in the parametric model.

In order to obtain the asymptotic value for the estimate of  $G_o$ , it is typically assumed that the structure of G is sufficiently complex so that the cost function achieves its minimal value for every causal transfer function G. Then, by splitting the

bias-pull into its causal and anti-causal parts, we have that, for BJ models, the estimate of  $G_o$  tends to

$$\hat{G} \to G_o + \frac{H_*}{M_u} \left[ (H_o - H_*) \frac{M_u}{H} \frac{\Phi_{wu}}{\Phi_u} \right]_+$$
(34)

where  $M_u$  is stable, minimum-phase and is such that  $\Phi_u = M_u M_u^*$ . This factorization is usually called a canonical factorization (see e.g. [18]).

The previous result has been shown in [8,23] for direct identification. We will see, in the sequel, that a similar result can also be obtained for the VCL method.

**Lemma 12** Under assumptions 1, 5, and 8, the cost function given in (15) and (13) is asymptotically (in the number of data points) given by:

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\epsilon} = \sigma_{w}^{2} + \frac{1}{2\pi} \int_{-\pi}^{\pi} \left| \frac{M}{K} \right|^{2} \Phi_{m} \left| X + \frac{\Phi_{nm}}{\Phi_{m}} \right|^{2} + \frac{1}{2\pi} \int_{-\pi}^{\pi} \left| \frac{M}{K} \right|^{2} \left[ \Phi_{n} - \frac{|\Phi_{nm}|^{2}}{\Phi_{m}} \right] \quad (35)$$

where

$$X = \frac{G_o - G}{F_1 + F_2 G}$$
(36)

$$m_t = u_t + \bar{C} \frac{H_o}{1 + \bar{C}G_o} w_t \tag{37}$$

$$n_t = \left[\frac{H_o}{F_1 + F_2 G_o} - \frac{K}{M}\right] w_t \tag{38}$$

**Proof:** From lemma 6, we have that the prediction error in (13) is given by:

$$_{t} = \bar{H}^{-1}[y_{t} - Gu_{t}] \tag{39}$$

$$=\bar{H}^{-1}[G_o u_t + H_o w_t - G u_t]$$
(40)

Hence,

 $\epsilon$ 

$$\epsilon_t = \frac{M}{K} \left[ \frac{G_o - G}{F_1 + F_2 G} u_t + \frac{H_o}{F_1 + F_2 G} w_t - \frac{K}{M} w_t \right] + w_t$$

$$= \eta_t + w_t$$
(41)
(42)

where

$$\eta_t = \frac{M}{K} \left[ X u_t + \left( \frac{H_o}{F_1 + F_2 G} - \frac{K}{M} \right) w_t \right]$$
(43)

$$=\frac{M}{K}\left[Xm_t + n_t\right] \tag{44}$$

 $<sup>^5</sup>$  Here and in the sequel all integrals are with respect to the variable  $\omega.$ 

Using assumption 8 we have that  $\eta_t$  depends on past values of  $w_t$ . Then, considering that  $w_t$  and  $\eta_t$  are independent, and

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\epsilon} = \sigma_w^2 + \frac{1}{2\pi} \int_{-\pi}^{\pi} \Phi_{\eta}$$
(45)

Finally, calculating the spectrum of  $\eta_t$  and completing squares we obtain (35).

We then have the following result:

**Theorem 13** Using Virtual Closed Loop identification for a Box-Jenkins model (i.e. when  $G(q, \rho)$  and  $K(q, \eta)$  are independently parametrized), the bias-pull for the estimates of  $G_o(q)$  is given by:

$$B_{G} := \begin{cases} (\lambda - 1) \left[ G_{o} + \bar{C}^{-1} \right] & \text{if } \bar{C} \neq 0 \\ F_{1} \frac{\Phi_{nm}}{\Phi_{m}} & \text{if } \bar{C} = 0 \end{cases}$$
(46)

where

$$\lambda = \frac{1}{1 - F_2 \frac{\Phi_{nm}}{\Phi_m}} \tag{47}$$

Moreover, for a general family of models for <sup>6</sup> G in the class of causal Box-Jenkins models, the asymptotic estimate,  $\hat{G}$ , of  $G_o$  tends to:

$$\hat{G} \to \begin{cases} G_o \bar{\lambda} - \bar{C}^{-1} (1 - \bar{\lambda}) & \text{if } \bar{C} \neq 0 \\ G_o + F_1 \frac{K_*}{MM_m} \left[ \frac{MM_m}{K_*} \frac{\Phi_{nm}}{\Phi_m} \right]_+ & \text{if } \bar{C} = 0 \end{cases}$$
(48)

where

$$\bar{\lambda} = \frac{1}{1 - F_2 \frac{K_*}{MM_m} \left[\frac{MM_m}{K_*} \frac{\Phi_{nm}}{\Phi_m}\right]_+}$$
(49)

 $K_* = K(q, \eta_*)$ , and  $M_m$  is the canonical factor of  $\Phi_m$ , i.e.  $\Phi_m = M_m M_m^*$ .

**Proof:** Splitting the integrand in the cost function (35) in terms of its causal and anti-causal parts we have that the part of the cost function that depends on X is given by:

$$\sigma_w^2 + \frac{1}{2\pi} \int_{\pi}^{\pi} \left| \frac{MM_m}{K_*} X + \left[ \frac{MM_m}{K_*} \frac{\Phi_{nm}}{\Phi_m} \right]_+ \right|^2 \tag{50}$$

Then, we have that the causal solution of the optimization problem is given by:

$$X_* = -\frac{K_*}{MM_m} \left[ \frac{MM_m}{K_*} \frac{\Phi_{nm}}{\Phi_m} \right]_+$$
(51)

Finally, using equation (36), and re-arranging terms we obtain the result.  $\hfill \Box$ 

**Remark 14** Notice that the expression for the estimate G in equation (48) is similar to the one obtained in nonparametric identification (see e.g. [19,36]). This is mainly due to the fact that, in non-parametric identification, the estimate of the transfer function has as many degrees of freedom as the one used in the definition of bias-pull in Definition 11.  $\nabla \nabla \nabla$ 

We next analyze the impact of the choice of the filters  $F_i$  ( $i = 1 \cdots 4$ ) on the variance of the estimates for  $G_o$ . We assume that there is no under-modelling, i.e. there exist  $\theta = \theta_o = \left[\rho_o^T \eta_o^T\right]^T$  such that  $R(\rho_o) = R_o$  and  $K(\eta_o) = K_o$ . This is a standard assumption to develop the accuracy analysis for the estimates.

**Remark 15** The inverse of the covariance matrix of the vector of parameters  $\hat{\theta}$  is given by:

$$P_{\theta}^{-1} = \begin{bmatrix} A & B \\ B^T & D \end{bmatrix}$$
(52)

where 7

$$A = \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|K_o|^2} \frac{\partial R}{\partial \rho} \frac{\partial R}{\partial \rho}^H \Phi_x$$
(53)

$$B = \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|K_o|^2} \frac{\partial R}{\partial \rho} \frac{\partial K}{\partial \eta}^H \Phi_{xw}$$
(54)

$$D = \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|K_o|^2} \frac{\partial K}{\partial \eta} \frac{\partial K}{\partial \eta}^H \sigma_w^2$$
(55)

The inverse of the covariance of  $\hat{\rho}$  is given by:

$$P_{\rho}^{-1} = A - BD^{-1}B^T$$
 (56)

In addition,  $P_{\rho}^{-1}$  is bounded as follows [4]:

- $P_{\rho}^{-1} \leq \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|K_o|^2} \frac{\partial R}{\partial \rho} \frac{\partial R}{\partial \rho}^H \Phi_x$ . Moreover, equality holds if and only if B = 0.
- $P_{\rho}^{-1} \geq \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|K_o|^2} \frac{\partial R}{\partial \rho} \frac{\partial R}{\partial \rho}^H \left[ \Phi_x \frac{|\Phi_{xw}|^2}{\sigma_w^2} \right]$ . Moreover, equality holds if and only if there exists a nonfrequency dependent matrix  $\Gamma$  such that  $\Gamma \frac{\partial K}{\partial \eta} = \frac{\partial R}{\partial \rho} \Phi_{xw}$ (almost everywhere in  $\omega$ ), where the derivatives are evaluated at  $\theta_o$ .  $\nabla \nabla \nabla$

The previous remmark is valid for linear and non-linear controllers, and also valid for finite number of parameters (see [4] for details).

<sup>&</sup>lt;sup>6</sup> For "general family of models", we mean that there are no system order constraints.

<sup>&</sup>lt;sup>7</sup> Here and in the sequel  $x^H$  denotes the conjugate transpose of x.

**Claim 16** In the case that the true controller,  $C_l$  is linear and equal to the virtual controller,  $\overline{C}$ , then the covariance of the parameters of  $\hat{G}$ , obtained using PEM in the VCL framework, is given by:

$$P_{\rho}^{-1}\{VCL\} = \frac{N}{2\pi\sigma_w^2} \int_{-\pi}^{\pi} \frac{1}{|H_o|^2} \frac{dG}{d\rho} \frac{dG}{d\rho}^H \Phi_u^r \qquad (57)$$

with

$$\Phi_u^r = |C_l S_o|^2 \Phi_r \tag{58}$$

Moreover, the covariance of the parameters of G, also satisfy the following inequality:

$$P_{\rho}^{-1}\{Direct\} \ge P_{\rho}^{-1}\{VCL\}$$
(59)

where  $P_{\rho}\{Direct\}$  is the covariance obtained when using direct identification. Moreover, equality holds in (59) if and only if there exists a non-frequency dependent matrix  $\Gamma$  such that  $\Gamma \frac{\partial H}{\partial \eta} = \frac{\partial G}{\partial \rho} \Phi_{uw}$  (almost everywhere in  $\omega$ ), where the derivatives are evaluated at  $\theta_o$ .

**Proof:** The first part follows from Remmark 15, the relationship of the signals in the virtual controller, and considering that, in the case that  $\overline{C} = C_l$ , the signal  $x_t$  and  $w_t$  are not correlated. The second part is obtained by using Remmark 15 for the case of direct identification ( $F_1 = F_4 = 1$ , and  $F_3 = F_2 = 0$ ) and equation (57).

**Remark 17** The previous lemma shows that most indirect identification methods provide estimates with the same covariance (provided that the true controller is linear). This lemma generalizes the results presented in [11] obtained by using the asymptotic in the number of parameters  $(n \rightarrow \infty)$  covariance formula (see [23]).  $\nabla \nabla \nabla$ 

# 7 Design of the filters

The analysis in the previous section provides expressions for bias and variance for estimates obtained by VCL for systems operating in closed loop using high order model.

Theorem 13 shows a frequency by frequency expression for the bias-pull. It also shows that the bias-pull can be reduced (or eliminated) if the cross-spectrum  $\Phi_{nm}$  is reduced (or equal to zero). This cross-spectrum can be calculated as the conjugate of the following expression:

$$\Phi_{mn} = \left[\frac{K_o - K_*}{M}\right]^* \left[\Phi_{uw} + \left(\frac{F_2 H_o}{F_1 + F_2 G_o}\right)\sigma_w^2\right]$$
(60)  
$$= \left[\frac{H_o - H_*}{F_1 + F_2 G_o}\right]^* \left[\Phi_{uw} + \left(\frac{F_2 H_o}{F_1 + F_2 G_o}\right)\sigma_w^2\right]$$
(61)

where  $H_*$  is defined as follows:

$$H_* := \frac{F_1 + F_2 G_o}{M} K_* \tag{62}$$

**Remark 18** Note the similarities between the term on the right hand side of (60) and the cross-spectrum  $\Phi_{xw}$  (See (25)).  $\nabla \nabla \nabla$ 

The bias-pull can be shaped by reducing any of the two terms above in a particular frequency range of interest. If  $\overline{C} = 0$ , then the second term in (60) is equal to  $\Phi_{uw}$  (which does not depend on the filters  $\{F_i\}_{i=1..4}$ ). Thus, if one chooses  $\overline{C} = 0$  then the only way to reduce the bias in the estimates of  $G_o$  is by choosing the filters such that  $\frac{1}{F_1F_4}[H_oF_4 - K_*] \approx 0$  in the frequency range of interest. This condition is similar to the "whitening" procedure. It requires knowledge of the noise model  $H_o$ . In the particular case of direct identification  $(F_1 = F_4 = 1 \text{ and } F_2 = F_3 = \overline{C} = 0)$  we have that the only way to have a small bias-pull is by having a good model for  $H_o$ .

Theorem 13 provides a basis for choosing suitable values for  $F_1$ ,  $F_2$ ,  $F_3$ ,  $F_4$  in order to reduce the bias due to undermodelling in the noise transfer function  $H_0$  and to the presence of a non-linear controller. In particular, we see that the asymptotic bias is small under *either* of the following two conditions

•  $H_o - H$  is small (i.e.  $K_o - K$  is small), •  $\overline{C} - C_l^a$  is small

Note that this holds on a frequency by frequency basis, so it suffices for  $\overline{C}$  to be near the linear approximation of the true controller when  $H_o - H$  is large or for  $H_o - H$  to be small when  $\overline{C}$  is a poor representation of the true controller.

On the other hand, the flexibility provided by the use of all the filters  $\{F_i\}_{i=1..4}$  allows us to reduce the bias-pull by minimizing both terms in (60) in the frequency range of interest. We see from Theorem 13 that, for the VCL method, the estimate of  $G_o$  is biased towards the negative inverse of the virtual controller. This bias will be small provided we can make  $\bar{\lambda}$  close to 1; i.e. make  $\frac{\Phi_{nm}}{\Phi_m}$  small. It is not surprising that a sufficient condition to have  $\Phi_{nm} \approx 0$  is the same condition that we found in section 5.2 in order to reduce the correlation between  $x_t$  and  $w_t$ . Moreover, we can ensure that  $\frac{\Phi_{nm}}{\Phi_m}$  is small (relative to 1) provided we choose  $\bar{C}$  "close to" the true controller even if the latter is nonlinear and / or ill-defined. In addition, the first term in (60) can also be minimized provided a model for  $G_o$  and  $H_o$  is available. Hence, it makes sense to choose  $F_1$ ,  $F_2$  such that  $\overline{C} = F_2/F_1$  is close to the true controller. For example, if the true controller is a linear controller incorporating antiwindup protection for input saturation, then  $\bar{C}$  could be chosen as the linear controller without anti-windup. Also note that the expressions for the bias-pull for the case when the



Fig. 4. System utilized in the numerical example.

"true" controller is linear can be easily obtained by using the results presented in Section 5.

From (7), it may be tempting to think that a good choice for  $F_3$ ,  $F_4$  would be such that  $F_1F_4 = F_2F_3$  since this removes all noise from (7). However, in this case,  $R_o = F_4F_3^{-1}$  i.e. we learn nothing about  $G_o$ . Thus, it is necessary to design the filters  $\{F_i\}_{i=1..4}$  such that M is different from zero in the frequency range of interest.

An alternative choice for  $F_3$ ,  $F_4$  would be to use a-priori estimates  $G_x$ ,  $H_x$  for  $G_o$ ,  $H_o$  to render  $K_o \approx 1$ . In this case, we might try using a fixed value for K (namely 1) in (13). In this case, the virtual closed loop scheme reduces to an output error method linking the measured variable  $z_t$  to the model output  $\hat{z}_t$ . Of course, based on Theorem 13, bias may result if  $\frac{F_1 + F_2 G_o}{F_1 F_4 - F_2 F_3}$  is significantly different from  $H_o$ in frequency ranges where  $\bar{C}$  is a poor approximation to the true controller.

On the other hand, it is well known that using filtered data improves the quality of estimates [23, chapter 7]. In fact, in [24] it is shown that for the identification of a particular continuous-time system it is necessary to filter the inputoutput signals in order to obtain"good" estimates. In addition, filtering the data is also used in "advanced" instrumental variables techniques such as refined instrumental variables (RIV) [12].

The general advice is that one should remove the "bad" data by filtering i.e. one should choose a frequency range of interest where the assumptions made in our model (e.g. system structure) hold (see [26,25,38,14,6]).

In addition, Claim 16 shows that whenever there is no-undermodeling (i.e. no bias) then the best choice for the filters are, in general, given by direct identification (i.e.  $F_1 = F_4 =$  1, and  $F_2 = F_3 = 0$ ). Note that there exists a class of systems where direct identification and the VCL (with  $\overline{C} = C$ ) provide estimates with the same accuracy (see [3] for an example). Of course, in the case of under-modeling, it seems natural to choose different values for the filters in order to deal with the usual bias-variance trade-off.

If a "good" linear model for the controller is available, but no model for  $G_o$  and  $H_o$  is available, then our general advice to design the filters  $\{F_i\}_{i=1..4}$  is as follows:

- Choose a filter L(q) that selects the frequency range of interest (where one believes that all the assumptions made hold).
- Choose filter  $F_1(q) = L(q)$ ,  $F_2(q) = L(q)\overline{C}(q)$ , where  $\overline{C}(q)$  is a good linear approximation of the true controller.
- Choose  $F_3(q) = 0$ , and  $F_4(q) = L(q)$ .

If good models for  $G_o(q)$  and  $H_o(q)$  are available ( $G_x$  and  $H_x$ ), then choose  $F_3(q)$  and  $F_4(q)$  to render  $K_o(q) \approx 1$  and such that they contain the filter L(q). If a model for  $G_o(q)$  and a "good" linear model for the controller are available, then choose the filters as in the Dual-Youla approach (see Table 2). Of course, the design of the filters  $\{F_i\}_{i=1..4}$  depends on the particular problem of interest, and one might design the filters based on different criteria.

# 8 A numerical example

Consider a system described by:

$$y_t = G_o(q)u_t + v_t \tag{63}$$

$$v_t = H_o(q)w_t \tag{64}$$



Fig. 5. Left side: Bode diagram of  $G_o$  (red-solid line) and  $G_o D$  (magenta-dashed line), and the filter L(q) (black-dash-dotted line). Right side: Bode diagram of Ho.

where

$$G_o(q) = \frac{b_1^o q^{-1}}{1 + a_1^o q^{-1}} \tag{65}$$

$$H_o(q) = \frac{0.3814 + 0.3103q^{-1}}{1 - 0.8571q^{-1} + 0.5488q^{-2}} \tag{66}$$

and  $a_1^o = -1.105$ ,  $b_1^o = 0.3155$ , and  $w_t$  is zero mean Gaussian white noise with variance  $\sigma_w^2$ .

Since the system is unstable, we perform the identification in closed loop using the following nominal control law:

$$u_t = C(q)[r_t - y_t] \tag{67}$$

where the reference signal is zero mean Gaussian white noise with variance  $\sigma_r^2 = 1$  and  $C(q) = \frac{0.3q^{-1}}{1-0.5q^{-1}}$ . However, we will assume that C(q) is implemented over a communication network. Our motivation for this choice is the observation that control over networks has become a topic of considerable research interest in the last few years (see e.g. [27]). In this area, the controller receives and sends signals through a network (see Figure 4). This means that the controller and the process can be located at distant points. A common drawback of this approach is that some data in the up-link (from plant to controller) or in the down-link (from controller to plant) may be missing due to packet loss. In this case, it is necessary to have "smarter" controllers and actuators. We will capture the idea of data-dropouts in our control law. In the case that a control signal  $u_t$  is missed, it is common to hold the previous value for the input signal. For illustrative purposes, we use instead an actuation given by the average of the previous 10 control actions. We assume that the missing data satisfies a Bernoulli distribution with probability of loss data P.

We test the PEM-direct, RIV and the VCL identification techniques for two scenarios:

- Nominal case: No under-modelling in G<sub>0</sub>, and there is no missing data (v<sub>t</sub> = 1, ∀t).
- Non-nominal case: Under-modelling in  $G_0$  i.e. the data is generated by a system given by  $G_o(q)D(q)$  where  $D(q) = \frac{0.5609+0.3038q^{-1}}{1-0.1353q^{-1}}$  and there is missing data P = 0.5.

In order to illustrate the impact of under-modelling of  $H_o$ we use the simplest noise model for the different techniques under study. For VCL we use an output error model (i.e. K = 1 for VCL), and for direct identification and RIV we use an ARX model (since OE models are not suitable for unstable processes). We use a model for G with the same structure as  $G_o$ .

For PEM-direct we use the following algorithm:

$$\hat{\rho} = \arg\min_{\rho} \sum_{t=1}^{N} (\varepsilon_t^F)^2 \tag{68}$$

where

$$F_t = L(q)\varepsilon_t \tag{69}$$

$$\varepsilon_t = A(q, \rho)y_t - B(q, \rho)u_t \tag{70}$$

and  $A(q, \rho) = 1 + a_1 q^{-1}$ ,  $B(q, \rho) = b_1 q^{-1}$ .

For the RIV method we use the following iterative procedure to estimate  $\rho_o$  [12]:

(1) Choose an initial value for  $\hat{\rho}$ .

ε

(2) Calculate the transfer function  $G(q, \hat{\rho})$ .



Fig. 6. Bode-magnitude plot of  $G_{o}$  (black-dashed line) and  $\hat{G}$  for different Monte-Carlo simulations for PEM-direct (left-hand plot, blue line), VCL (plot in the middle, green line), and RIV (right hand plot, magenta) in non-nominal conditions using N = 100 data points.

- (3) Calculate the transfer functions  $T(q) = \frac{GC}{1+GC}$  and S(q) = 1 - T(q).

- (4) Calculate the signals  $y_t^r = T(q)r_t$  and  $u_t^r = S(q)r_t$ . (5) Calculate the vectors  $\varphi_t = [L(q)u_{t-1} L(q)y_{t-1}]^T$ and  $z_t^k = [L(q)u_{t-1}^r L(q)y_{t-1}^r]^T$ . (6)  $\hat{\rho}_k = [\sum z_t^k \varphi_t^T]^{-1} \sum z_t^k L(q)y_t$ . (7) Set  $\rho = \hat{\rho}_k$  go to step (2) until convergence.

For the VCL method we choose  $F_1(q) = L(q)$ ,  $F_2 =$  $0.6L(q), F_3(q) = 0, F_4(q) = L(q)$ . Thus, we have that  $\bar{C} = 0.6$  which is the dc-gain of the linear controller in the loop when there is no missing data.

For all of the techniques we choose the filter L(q) as a Butterworth filter of third order and having a cut-off frequency  $0.4/\pi [1/s]$ . We analyze the performance of all algorithms for N = 100 and N = 10000 data points by using 100 Monte-Carlo experiments for different values of  $\gamma$  given by

$$\gamma = \frac{\hat{\sigma}_v^2}{\hat{\sigma}_y^2} \tag{71}$$

where  $\hat{\sigma}_v^2$  and  $\sigma_{\bar{y}}^2$  are estimates of the noise and noise-free-output  $(\bar{y}_t = y_t - v_t)$  variance obtained from the data.

Figure 6 shows the Bode-magnitude diagram for the estimates of 100 Monte-Carlo runs for different values of  $\gamma$  obtained in the non-nominal case using N = 100. We see that the VCL method provides the most accurate estimates.

We calculate the average 2-norm of the relative error for the models obtained given by:

$$R_{e} = \frac{1}{2\pi N_{s}} \sum_{k=1}^{N_{s}} \int_{-\pi}^{\pi} \left| \frac{\hat{G}_{k} - G_{o}}{G_{o}} \right|^{2} d\omega$$
(72)

where  $N_s$  is the number of Monte-Carlo simulations, and  $\hat{G}_k$ is the corresponding model obtained from the data in each experiment. This is plotted in Figure 7 as a function of  $\gamma$ .

We see from Figure 7 that PEM-direct method provides good estimates when the noise is small. However, the quality of the estimates deteriorates when the noise is large.

We see that, under nominal conditions and large data-length, the RIV technique provides good estimates. However, in the non-nominal case and for short data-length the estimates are not satisfactory.

On the other hand, the estimates obtained by VCL are good in both scenarios irrespective of data-length.

#### 9 Conclusions

In this paper we have presented a general method (the Virtual Closed Loop method) to perform identification of systems operating in closed loop. The method is sufficiently



Fig. 7. Relative error  $R_e$  for different values of  $\gamma$  under nominal conditions (N = 100: top-left plot, N = 10000: top-right plot) and non-nominal conditions (N = 100: bottom-left plot, N = 10000: bottom-right plot) conditions. VCL (black-solid line), PEM-direct (blue-dashed line), RIV (magenta-dash-dotted line).

general to be applied to different types of systems. We have presented a correlation analysis of the signals of interest, analyzed asymptotic bias due to feedback and noise model mismatching, and also the impact on the variance of  $\hat{G}$  at different frequencies. We have shown that the new parametrization generalizes known methods for closed loop identification and also offers additional flexibility. A numerical example has confirmed the claimed merits of the approach.

# References

- J. C. Agüero. System identification methodologies incorporating constraints. PhD thesis, School of Electrical Engineering and Computer Science, The University of Newcastle, Australia, 2005.
- [2] J. C. Agüero and G. C. Goodwin. Virtual closed loop identification: A Subspace Approach. In *Proceedings of the 43rd IEEE Conference* on Decision and Control, CDC, pages 364–369, Atlantis, Paradise Island, Bahamas, 2004.
- [3] J. C. Agüero and G. C. Goodwin. On the optimality of open and closed loop experiments in system identification. In 45th IEEE Conference on Decision and Control (CDC), San Diego, California, 2006.
- [4] J. C. Agüero and G. C. Goodwin. Choosing between open and closed loop experiments in linear system identification. *IEEE Trans. Automatic Control*, 52(8):1475–1480, 2007.
- [5] J. C. Agüero, G. C. Goodwin, and P. M. J. Van den Hof. Virtual closed loop identification: A generalized tool for identification in closed loop. In *Proceedings of the 47th IEEE Conference on Decision* and Control, CDC, 2008.
- [6] J. C. Agüero, J. I. Yuz, G. C. Goodwin, and R. A. Delgado. On the equivalence of time and frequency domain maximum likelihood estimation. *Automatica*, 46(2):260–270, 2010.
- [7] U. Forssell. Closed-loop Identification. Methods, theory, and applications. PhD thesis, Linköping University, Linköping, 1999.

- [8] U. Forssell and L. Ljung. Closed loop identification revisited. *Automatica*, 35:1215–1241, 1999.
- [9] U. Forssell and L. Ljung. A projection method for closed loop identification. *IEEE Trans. Automatic Control*, 45(11):2101–2106, 2000.
- [10] M. Gevers. Identification for control: from the early achievements to the revival of experiment design. *European Journal of Control*, 11(4-5):335–352, 2005.
- [11] M. Gevers, L. Ljung, and P. M. J. Van den Hof. Asymptotic variance expressions for closed-loop identification. *Automatica*, 37(5):781– 786, 2001.
- [12] M. Gilson, H. Garnier, P. Young, and P. M. J. Van den Hof. Refined instrumental variable method for closed-loop system identification. In 15th IFAC Symposium on System Identification, Saint Malo, France, 2009.
- [13] M. Gilson and P. M. J. Van den Hof. Instrumental variable methods for closed-loop system identification. *Automatica*, 41(2):241–249, 2005.
- [14] G. C. Goodwin, J. C. Agüero, J. S. Welsh, J. I. Yuz, G. J. Adams, and C. R. Rojas. Robust identification of process models from plant data. *Journal of Process Control*, 18:810–820, 2008.
- [15] G. C. Goodwin, S. F. Graebe, and M. E. Salgado. *Control System Design.* Prentice Hall, Upper Saddle River, NJ, 2001.
- [16] G. C. Goodwin and R. Payne. Dynamic System Identification: Experiment design and data analysis. Academic Press, 1977.
- [17] F. Hansen, G. Franklin, and R. Kosut. Closed loop identification via the fractional representation: Experiment design. *Proceedings of the American Control Conference*, pages 1422–1427, 1989.
- [18] B. Hassibi, A. H. Sayed, and T. Kailath. Indefinite quadratic estimation and control. A unified approach to  $H^2$  and  $H^{\infty}$  theories. SIAM, 1999.
- [19] WP Heath. Bias of indirect non-parametric transfer function estimates for plants in closed loop. *Automatica*, 37(10):1529–1540, Jan 2001.
- [20] H. Hjalmarsson. From experiment design to closed-loop control. Automatica, 41(3):393–438, 2005.

- [21] H. Hjalmarsson, M. Gevers, and F. De Bruyne. For model-based control design, closed-loop identification gives better performance. *Automatica*, 32(12):1659–1673, 1996.
- [22] L. Ljung. Convergence analysis of parametric identification methods. *IEEE Transactions on Automatic Control*, AC-23:770–783, 1978.
- [23] L. Ljung. *System Identification: Theory for the user*. Prentice Hall, 2nd edition, 1999.
- [24] L. Ljung. Aspects and experiences of user choices in subspace identification methods. In 13th IFAC Symposium on System Identification, pages 1802–1807, Rotterdam, The Netherlands, 2003.
- [25] T. McKelvey. Frequency domain identification methods. *Circuits systems signal processing*, 21(1):39–55, 2002.
- [26] R. Pintelon and J. Schoukens. Box-jenkins identification revisitedpart i: Theory. Automatica, 42(1):63–75, 2006.
- [27] L. Schenato, B. Sinopoli, M. Franceschetti, K. Poolla, and S. Sastry. Foundations of control and estimation over lossy networks. *Proceedings of the IEEE*, 95(1):163 – 187, 2007.
- [28] R. J. P. Schrama. An open-loop solution to the approximate closedloop identification problem. In *Identification and System Parameter Estimation 1991*, pages 761–766, Budapest, Hungary, 1991.
- [29] T. Söderström. *Discrete-Time Stochastic Systems: Estimation and control*. Prentice Hall, second edition, 2002.
- [30] T. Söderström and P. Stoica. System identification. Prentice-Hall International, 1989.
- [31] T. Söderström, P. Stoica, and E. Trulsson. Instrumental variable methods for closed-loop systems. In *Proc. 10th IFAC World Congress*, pages 363–368, Munich, Germany, 1987.
- [32] A. Stuart, J. K. Ord, and S. Arnold. *Kendall's Advanced Theory of Statistics*, volume 2A. Edward Arnold, 1999.
- [33] P. M. J. Van den Hof. Closed-loop issues in system identification. Annual Reviews in Control, 22:173–186, 1998.
- [34] P. M. J. Van den Hof and R. J. P. Schrama. Identification and control–closed loop issues. *Automatica*, 31(12):1751–1770, 1995.
- [35] E. T. Van Donkelaar and P. M. J. Van den Hof. Analysis of closedloop identification with a tailor-made parameterization. *European Journal of Control*, 6(1):54–62, 2000.
- [36] J. S. Welsh and G. C. Goodwin. Finite sample properties of indirect nonparametric closed-loop identification. *IEEE Transactions* on Automatic Control, 47(8):1277–1292, 2002.
- [37] H. White. *Estimation, Inference and Specification Analysis.* Cambridge University Press, 1996.
- [38] J.I. Yuz and G.C. Goodwin. Robust identification of continuous-time systems from sampled data. In H. Garnier and L. Wang, editors, *Continuous-time Model Identification from Sampled Data*, chapter 3, pages 67–89. Springer, 2008.
- [39] W.X. Zheng and C.B. Feng. A bias-correction method for indirect identification of closed-loop systems. *Automatica*, 31(7):1019–1024, 1995.